

Rotational evolution of the Vela pulsar during the 2016 glitch

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The 2016 Vela glitch observed by the Mt Pleasant radio telescope provides the first opportunity to study pulse-to-pulse dynamics of a pulsar glitch, opening up new possibilities to study the neutron star’s interior. We fit models of the star’s rotation frequency to the pulsar data, and present three new results. First, we constrain the glitch rise time to less than 12.6 s with 90% confidence, almost three times shorter than the previous best constraint. Second, we find definitive evidence for a rotational-frequency overshoot and fast relaxation following the glitch. Third, we find evidence for a slow-down of the star’s rotation immediately prior to the glitch. The overshoot is predicted theoretically by some models; we discuss implications of the glitch rise and overshoot decay times on internal neutron-star physics. The slow down preceding the glitch is unexpected; we propose the slow-down may trigger the glitch by causing a critical lag between crustal superfluid and the crust.

Pulsar glitches, rotational irregularities of otherwise stably rotating neutron stars, are believed to be caused by the complex interplay between micro- and macrophysical properties of the star’s internal components. One model posits that superfluid vortices in the inner crust suddenly unpin, transferring angular momentum to the star’s lattice crust¹. This is seen as an increase in the frequency of pulsations. Such models are difficult to verify; the internal components of the star are shielded from view and their behavior has to be inferred indirectly. Until recently, radio observations of glitches were limited to observations before and after the glitch, but not during. The detailed morphology of glitch dynamics (e.g., the glitch rise time) is therefore not well constrained or understood. In 2016, the first pulse-to-pulse observations of a glitch were made using the University of Tasmania Mt Pleasant 26 m radio telescope². Those observations showed variations in pulse shape of four pulses starting 20 rotations before the inferred time of the glitch, which are attributed to variations in the magnetospheric state. A preliminary estimate of ~ 4.4 s for the glitch rise time was given.

In this Article, we provide a detailed pulse-to-pulse analysis of the glitch morphology. Our main results are three-fold. First, we constrain the glitch rise time to less than 12.6 s with 90% confidence. We connect this with internal neutron-star physics using a body-averaged description of the components participating in the glitch. Second, we show that a frequency overshoot—an increase in the rotation frequency above the post-glitch equilibrium value—and subsequent fast relaxation exist immediately following the glitch, in agreement with the only two previous high time-resolution observations of Vela glitches^{3,4}, and as explained by several models of neutron-star glitches^{5–8}. Third, we show that the glitch may be preceded by an initial precursor slow-down, whereby the crust of the star slowed before rapidly speeding up; we speculate that this preceding slow-down of the pulsar’s rotation triggered the glitch.

In the superfluid vortex model of pulsar glitches, the star’s crust, and components tightly coupled to it, spin down due to external electromagnetic torques. The crustal superfluid is decoupled from the lat-

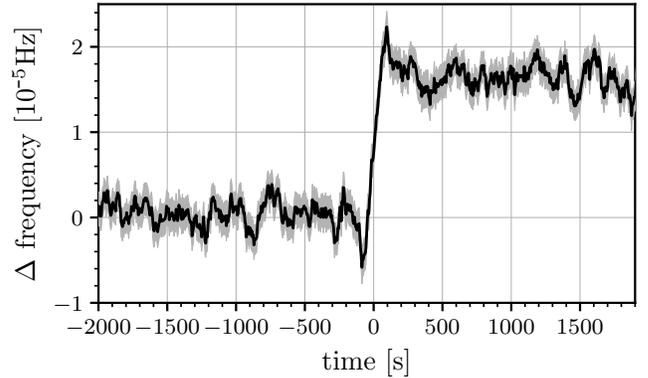


Figure 1 | Rotational frequency evolution: we fit a constant-frequency model in 200 s-long sliding windows. The frequency and time are given relative to nominal values; see text for details. The sliding window elongates features with respect to their true temporal evolution, and suppresses stochastic fluctuations of the frequency. Despite this, the largest fluctuations can be seen immediately following the glitch (the overshoot) and immediately prior to the glitch (the precursor slow-down).

tice as a result of vortex pinning, implying a lag develops between the superfluid’s angular velocity Ω_{sf} and the crust’s angular velocity Ω_{crust} ; $\Omega_{\text{lag}} \equiv \Omega_{\text{sf}} - \Omega_{\text{crust}} > 0$. When Ω_{lag} reaches a critical value Ω_{crit} , some mechanism initiates the glitch, simultaneously releasing a large number of vortices. The superfluid’s excess angular momentum is transferred to the crust, producing the observed spin up. The specific trigger for the angular-momentum transfer is not well understood, although speculation abounds^{9–16}.

We show there exist fluctuations of Vela’s spin period prior to the glitch, see Fig. 1, and speculate that $\Omega_{\text{lag}} > \Omega_{\text{crit}}$ is reached due to a single, relatively large stochastic fluctuation. This is independent of the glitch trigger mechanism, but provides a means for reaching the critical lag between superfluid and rigid crust; we discuss implications below.

1 Frequency evolution

For a model-agnostic view of the pulsar’s evolution, we fit^{17–19} a constant-frequency model to 200 s-long data segments, sliding this window throughout the data (for details, see the Methods section). In Fig. 1 we show the median frequency and 90% credible interval for each window. The frequency evolution on the vertical axis is given as the difference between the inferred frequency and a nominal value of 11.186433 Hz. Times are relative to the solar-system barycentre time of the fitted glitch MJD 57734.4849906². As we detail below, this analysis method allows for a powerful comparison with physical models of pulsars’ rotational evolution. The glitch can clearly be seen near time zero. We caution that one cannot use this plot to find the glitch rise time as the sliding window time averages the frequency evolution, thereby elongating such features. The time window also suppresses the amplitude of features whose timescale is shorter than the window length. Despite this, Fig. 1 shows stochastic fluctuations in the frequency evo-

model	$f(t)$ parameters	$\log_{10} B$
H_s	$\Delta f_r = \Delta f_d = 0$	—
H_1	$\Delta f_r \neq 0, \Delta f_d = 0$	-1.7
H_2	$\{\Delta f_r, \Delta f_d\} \neq 0$	0.38
H_{2+p}	see Eq. (3)	2.8

Table 1 | Definitions and Bayes factors for the four primary models tested in this Article, see Eq. (1) and the consistency relation. The final column is the log-Bayes factor between each model and the step-glitch model H_s . Uncertainties on the log-Bayes factors are $\lesssim 0.2$.

lution for the ~ 2000 s before and after the glitch. The largest fluctuations are immediately before, and immediately after the glitch. This motivates the more thorough analysis detailed below.

To quantitatively analyse the frequency evolution, we define three models. The most general of these is H_2 , where rotational changes are characterised by a constant term plus *two* exponentials,

$$f_2(t) = f_0 + H(t-t_g) \left[\Delta f + \Delta f_r e^{-\frac{(t-t_g)}{\tau_r}} + \Delta f_d e^{-\frac{(t-t_g)}{\tau_d}} \right], \quad (1)$$

where $H(t-t_g)$ is the Heaviside step function, t_g is the glitch time, and $\Delta f, \Delta f_r, \Delta f_d$ are the glitch magnitude and amplitudes of each exponential term. Equation (1) is supplemented by the relation $\Delta f + \Delta f_r + \Delta f_d = 0$ ensuring the frequency evolution is continuous at the glitch. This relation, along with positive log-uniform priors on Δf and Δf_d imply $\Delta f_r \leq 0$; in Eq. (1), the first exponential therefore describes the “rise” in the frequency on time-scale τ_r , while the second exponential describes a “decay” in frequency on time-scale τ_d .

Models H_s and H_1 , specified in Tab. 1, are limiting cases of H_2 . These phenomenological models are motivated by analytic solutions to coupled rigid-body problems, discussed later where we also introduce a final model H_{2+p} .

For each model, we integrate the frequency evolution to obtain the phase evolution, which we invert to obtain the model-predicted arrival time for each pulse. A likelihood of the model given the data is calculated by modelling the pulse arrival time as a sum of the deterministic arrival time predicted by the timing model, and a zero-mean Gaussian process with unknown variance. Using this likelihood and a suitable set of priors for the model parameters, we infer their posterior distributions and the evidence for the model using PyMultiNest^{20–22}. Complete descriptions of the likelihood and prior are given in the Methods.

The simplest model in Tab. 1, H_s —*step glitch*, ignores the complex morphology of the glitch and models the frequency evolution as a simple step function of amplitude Δf at time t_g . We use this as a base-model against which we compare all other models in Tab. 1. For the H_s model, we infer a glitch magnitude $\Delta f = 16.11^{+0.04}_{-0.04}$ μHz and time, $t_g = -0.31^{+2.74}_{-2.78}$ s (given relative to the reported value) consistent with the initial observation².

2 Glitch rise time

We analyse the rise time using a simple, physically-motivated reference model; a body-averaged model with two uncoupled spinning components that suddenly couple, has equations of motion that can be integrated to give a model with a single exponential rise time τ_r ²³, corresponding to model H_1 in Tab. 1. In the limit where τ_r is much smaller than the pulse period, the H_1 model is equivalent to the step-glitch model H_s .

In Fig. 2, we show the τ_r posterior for the H_1 model. The posterior peaks at zero; we cannot resolve the rise time of the glitch. This is consistent with the Bayes factor between H_1 and H_s being in favour of the simpler step-glitch model; Tab. 1. Nevertheless, the τ_r posterior gives

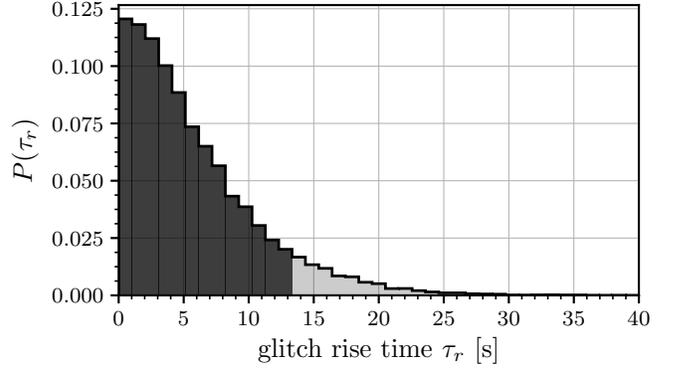


Figure 2 | Posterior distribution $p(\tau_r)$ for the glitch rise time, τ_r . The dark region marks the 90% confidence interval $\tau_r \leq 12.6$ s. The distribution peaks at zero, consistent with the Bayes factor which supports the simpler H_s model.

a 90% upper limit of $\tau_r \leq 12.6$ s. This improves upon the previous best upper limit on the rise time of the 2004 Vela glitch of $\tau_r \lesssim 30$ s⁴.

We also fit a model with a logistic function such that $f(t) \propto 1/(1 + e^{-t/\tau_r})$. While the functional form differs, it captures a similar idea of a rise in frequency. Results are similar to the H_1 model; the Bayes factor favours the step glitch, with an upper limit $\tau_r \leq 7.64$ s. The evidence for the logistic and H_1 models are comparable, implying neither is substantially favoured by the data. The physically-motivated reference model H_1 and the logistic model are phenomenological; we expect the true evolution of the glitch rise to be more complicated, although we know of no robust predictions in the literature. Nevertheless, the use of these two models shows our derived rise time is relatively insensitive to the details of the mathematical model. Throughout this work, we quote the more conservative upper-limit rise time of the H_1 model.

Within body-averaged models, the glitch rise is described by a dimensionless mutual-friction coefficient \mathcal{B} , controlled by the underlying vortex dynamics^{8,24–26}. Invoking a simplified two-component model, where the superfluid in the inner crust provides the angular-momentum reservoir for the glitch, the rise time is an indirect measurement of the coupling strength between the superfluid and the crust, with moments of inertia I_{sf} and I_{crust} , respectively. We derive a lower limit

$$\mathcal{B} \gtrsim 5.7 \times 10^{-6} \left(\frac{\tau_r}{12.6 \text{ s}} \right)^{-1} \left(\frac{f_{\text{sf}}}{11 \text{ Hz}} \right)^{-1} \left(\frac{I_{\text{sf}}/I_{\text{tot}}}{0.01} \right), \quad (2)$$

where f_{sf} is the rotation frequency of the superfluid and $I_{\text{tot}} = I_{\text{sf}} + I_{\text{crust}}$.

3 Glitch overshoot and relaxation

To investigate the overshoot and subsequent relaxation, we include a second exponential in the frequency evolution. This H_2 model (Tab. 1) is a simplified version of the three-component neutron-star model from Ref.⁸, where the star is separated into crustal superfluid, core superfluid and the non-superfluid crust component. The H_2 model assumes the three constituents are rigidly rotating and coupled via constant mutual-friction coefficients. While the specific equation chosen to model the overshoot and decay is motivated by this three-component model, we treat it as phenomenological for understanding the glitch dynamics. This phenomenological model could also be interpreted in terms of alternative physical models that also predict frequency overshoots^{5–7}.

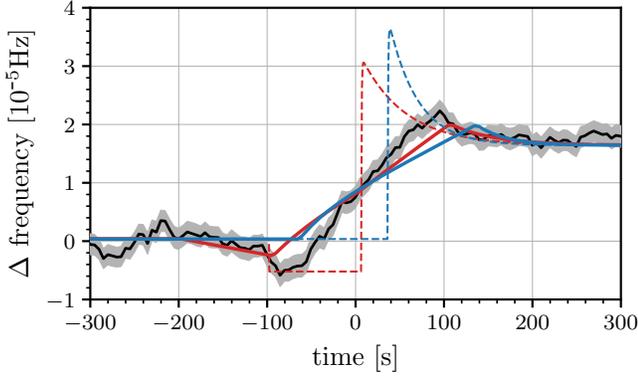


Figure 3 | Rotational frequency evolution of the data (black and grey; reproduced from Fig. 1) and best-fit models. We show the maximum-likelihood fit for the model that includes a step glitch with an overshoot and subsequent exponential decay (H_2 ; blue curves), and for a model that includes a slow-down preceding the glitch (H_{2+p} ; red curves). Dashed curves show the raw frequency evolution. Solid curves shows the time-averaged frequency evolution, which can be compared directly with the time-averaged data (black).

We fit model H_2 to the data, and show the maximum likelihood alongside the time-windowed data in Fig. 3. We show the raw model (dashed blue curve) and the time-averaged frequency evolution (solid blue curve); the latter can be directly compared to the time-windowed data (black curve).

Comparing the overshoot-decay model H_2 and step-glitch model H_s yields a Bayes factor $\log_{10} B = 0.38$, providing marginal support for the overshoot model. However, the H_1 model (which compared unfavourably against H_s) is a special case of model H_2 . The more relevant Bayes factor to understand the importance of the overshoot and relaxation is between H_2 and H_1 , for which $\log_{10} B = 2.1$ showing substantial evidence in favour of the overshoot and relaxation. Alternatively, we can compare H_s with a modified step-function evolution including a single decaying exponential. This model, often used in glitch-timing to model long-term, $\mathcal{O}(\gtrsim 1 \text{ day})$, relaxation was also fit to the data: the Bayes factor, $\log_{10} B = 2.0$, demonstrates strong support for a relaxation component with $\tau_d \sim 1 \text{ min}$, further confirming the existence of the overshoot. We remind the reader that a Bayes factor of $\log_{10} B > 2$ is considered “decisive” support for a model, while $1 \leq \log_{10} B \leq 2$ is considered “strong” support²⁷.

The existence of the overshoot is clear both visually and through our quantitative analysis. This is *not* the first identification of an overshoot: Refs.^{3,4} found a similar feature in the 2000 and 2004 Vela glitches, the only other pulsar glitches with high-time resolution data. However, these were not as well resolved as the telescope was less sensitive, requiring 10 s-folding of the pulses to achieve sufficient signal-to-noise ratio to calculate times of arrival. Comparing with the work herein, the pulse folding also likely explains the less constrained $\lesssim 30 \text{ s}$ glitch rise time.

The maximum-likelihood overshoot-decay model H_2 shown in Fig. 3 has a decay timescale of $\tau_d = 65.97_{-24.44}^{+59.38} \text{ s}$, magnitude $\Delta f_d = 9.36_{-6.38}^{+12.20} \mu\text{Hz}$, and the rise time is similarly constrained as in the H_1 model. The large uncertainty on the decay time is due to a strong correlation with the size of the overshoot: larger overshoots with shorter decay times are just as probable as smaller overshoots with longer decay times. We discuss this in more detail below, including Fig. 4, which shows the covariance between the size of the overshoot and the decay timescale using model H_{2+p} .

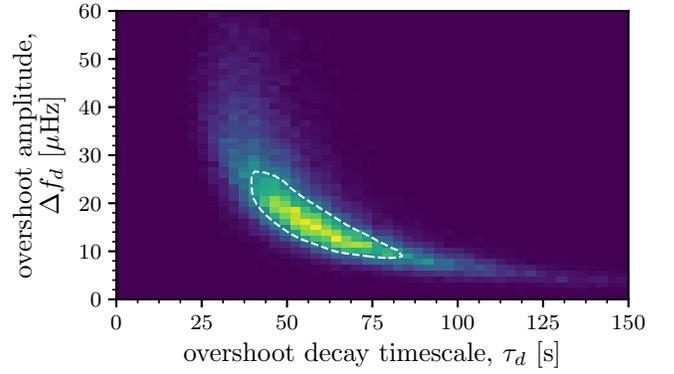


Figure 4 | Posterior for overshoot-decay parameters in the H_{2+p} model, corresponding to the red curves in Fig. 3. We show the overshoot decay timescale τ_d and frequency amplitude of the overshoot Δf_d . The white dashed contour indicates the one-sigma confidence level.

4 Slow-down preceding the glitch

We investigate the frequency slow-down preceding the glitch by extending the H_2 model to include a step-function in frequency before the glitch. This phenomenologically models a spin-down event sometime before the glitch. The frequency evolution for H_{2+p} , the *precursor slow-down model*, is given by

$$f_p(t) = f_2(t) - \Delta f_p \Pi(t; t_g - \Delta t, t_g), \quad (3)$$

where $f_2(t)$ is given by Eq. (1), Π is a rectangle function such that the frequency decreases by Δf_p for the period Δt prior to the glitch. Constructing the model in this way, Δf remains the long-term frequency change at the glitch.

Model H_{2+p} is phenomenological and motivated by the data; Fig. 1. Furthermore, it is one of many simple phenomenological models that could be used; e.g., exponential or linear drift models. Rather than perform a systematic study, we focus solely on H_{2+p} with the aim to motivate further research in this area. To this end, we speculate below about the causes of the slow-down. Until we have a more physically-grounded model, the significance of the slow-down is difficult to establish.

The Bayes factor shows the precursor slow-down model H_{2+p} is the preferred of all models tested here: comparing with the overshoot-decay H_2 model (the next most preferred), $\log_{10} B = 2.5$. This suggests the data supports a slow down of the rotation prior to the glitch, in addition to an overshoot and decay. In Fig. 3, we show the best-fit slow-down model in red; the dashed curve represents the raw frequency evolution, and the solid curve the time-averaged best-fit model.

In Fig. 4, we show the posterior for the size of the overshoot Δf_d and the overshoot relaxation timescale τ_d using the H_{2+p} model. As mentioned, these are inversely correlated, implying a wide range of equally-likely values for both parameters.

In Fig. 5, we show the posterior of the precursor slow-down Δf_p and the time before the glitch at which this occurs Δt . Although the size of the slow-down is not well constrained with $\Delta f_p = 5.40_{-2.05}^{+3.39} \mu\text{Hz}$, we note this is a significant fraction of the actual glitch size $\Delta f = 16.01_{-0.05}^{+0.05} \mu\text{Hz}$ for the H_{2+p} model.

We use a half-normal prior distribution on Δf_p (see Table 2, *mentary Material*), which places the maximum prior probability at zero, gives reasonable support over values $\lesssim 10^{-5} \text{ Hz}$, and exponentially disfavours larger positive values. If, on the other hand, we use uniform priors, then another local maxima in the posterior distribution becomes

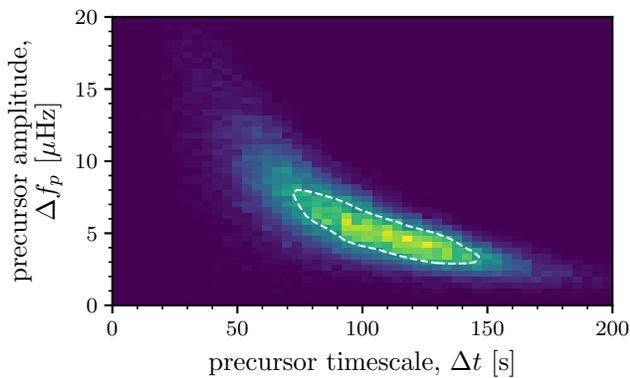


Figure 5 | Posterior for precursor parameters in the H_{2+p} model, corresponding to the red curves in Fig. 3. We show the precursor time Δt and amplitude of the frequency slow-down Δf_p . The white dashed contour indicates the one-sigma confidence level.

present at significantly larger Δf_p and at a shorter time preceding the glitch. We find it difficult to physically motivate a precursor slow-down many times larger than the glitch size; this explains our choice of the half-normal prior. However, as all our models are phenomenological, we leave open the possibility that this larger mode exists, and leave that for future exploration.

The Vela pulsar has been seen glitching three times with high-time resolution observations²⁻⁴. The first two observations were with a less-sensitive instrument, implying pulses must be folded to obtain sufficient signal-to-noise. The method detailed herein can be used on those data, and we encourage reanalysis of that data in search of precursor slowdowns and overshoot-decays; such features have been hinted at in Ref.³.

We divide models of the physical nature of the precursor slowdown into two groups. In the first, the fluctuation seen prior to the glitch in Fig. 1 is a large incarnation of the stochastic fluctuations seen preceding and following the glitch. In the second group, the slow-down is unrelated to this stochastic noise.

In future work, we will probe the hypothesis that the stochastic fluctuations cause the glitch by developing a statistical analysis of the Vela noise using data away from the glitch. This requires development of a spin-down model in the absence of a glitch that takes into account the stochastic fluctuations, with subsequent analysis of a large number of off-glitch data segments. Comparing the amplitude of the frequency slow-down ($\sim 5 \mu\text{Hz}$) with the noise distribution, could indicate if the slow-down is a statistical outlier from the typical noise, potentially allowing falsification of the idea that the spin-down event caused the glitch.

The cause of the stochastic fluctuations in Fig. 1 are also of interest. While they could be a manifestation of jitter noise, they may be due to fluctuations in the rotation rate caused by instabilities, or to extrinsic effects such as fluctuations in dispersion measure or scattering²⁸. These ideas can be explored by looking at other sets of high time-resolution data, using the same method used to produce Fig. 1.

We hypothesise that the slow-down may be due to some intrinsic mechanism in the star, although this is far from certain. Reference² reported short-timescale variations in pulse shape during the glitch, including a null pulse and unusual pulse shapes before and after the null. These could indicate the glitch or preceding slow-down are magnetospheric in origin, although it is difficult to disentangle cause and effect. Moreover, the large change in spin period on such short timescales is difficult to quantitatively explain without catastrophically changing the

magnetic-field topology, which would likely be accompanied with a long-term change in pulse shape as observed following glitches in high magnetic-field pulsars and magnetars²⁹⁻³¹. Such a long-term change in pulse shape has not been observed following the 2016 Vela glitch².

5 Did the spin-down event cause the glitch?

That the glitch is preceded by a spin-down event is intriguing. Many models^{6,8,32-35} posit that glitches are triggered when sufficient lag is built up between the superfluid component of the inner crust and the lattice crust. Some fraction of the fluctuations in Fig. 1 may be due to intrinsic, stochastic variations of the star's rotational period. Those variations may take place on timescales faster than the coupling timescales of the crust and internal components. If the slow-down event is one such stochastic variation, albeit a large one, we hypothesise this may trigger the glitch by spinning down the crust and driving the lag above its critical value.

Our hypothesis has natural corollaries for glitch statistics of the pulsar population: if the stochastic variations are large or comparable to the change in spin period from dipole radiation, glitches will occur probabilistically when the combination of the spin down and variations takes the crust-core lag above the trigger threshold. In such cases, the time period between glitches would neither be regular nor Poisson-distributed, but would depend on the relative size of the variations with respect to the spin down. When the variations are large with respect to the spin down, the glitch recurrence time should be Poisson-distributed. Finally, if the variations are small compared to changes from magnetic spin down, the glitch recurrence time should only depend on the spin-down timescale of the system. Additionally, we expect distinct behavior for those stars where glitches are driving the system far from the critical point and spin down is required to return to the threshold before variations can initiate a subsequent glitch, versus the objects where this is not the case and variations can always trigger a glitch.

It is worth noting that, in reality, the critical lag will depend on the neutron-star density and not have a single value, but there is likely also a statistical distribution associated with the macroscopically-averaged value of the critical lag. This should not significantly effect the arguments presented above as they are mainly qualitative. However it should be taken into account when quantitatively defending this model.

Although there are two populations of pulsars according to their glitch recurrence statistics³⁶⁻³⁸, more work is required to establish whether the two populations correlate with the relative magnitude of the pulsar's stochastic variations and their spin-down timescales. Such a statistically rigorous study would potentially be difficult; it is not clear which fluctuations are related to intrinsic pulsar spin noise and, e.g., pulse-jitter noise. This would also be difficult to generalize to other pulsars.

6 Conclusion

During the 2016 glitch, the Vela pulsar first spun down. A few seconds later it rapidly spun up, before finally spinning down with an exponential relaxation time of ~ 60 s. This model is substantially favoured over a simple step glitch, or one with only a single spin-up event (see Tab. 1).

Testing the rise time alone, we constrain $\tau_r \leq 12.6$ s (90% confidence), consistent with the estimated value of Ref.², and reducing the previous-best constraint of $\tau_r \lesssim 30$ s for the 2004 Vela glitch⁴. Invoking a two-component neutron-star model, our new constraint translates into a lower limit for the mutual-friction coupling of $\mathcal{B} \gtrsim 5.7 \times 10^{-6}$; Eq. (2).

We find a frequency overshoot and exponential relaxation with amplitude $\Delta f_d = 17.77^{+13.68}_{-7.99} \mu\text{Hz}$ and decay time scale $\tau_d = 53.96^{+24.02}_{-14.82}$ s for the H_{2+p} model, a feature that can be theoretically

explained^{5–8}. For example, within the three-component model of Ref.⁸, the overshoot only exists if the crust mutual-friction coefficient (coupling the crustal superfluid and crust) exceeds the core friction coefficient (coupling the core superfluid and crust). Providing a more qualitative analysis of the internal physics, e.g., constraining moments of inertia and coupling coefficients, is difficult at this point as the phenomenological models studied herein do not produce sufficient information.

Finally, we find evidence for a slow-down, or possibly a precursor *antiglitch*, immediately before the glitch. To the best of our knowledge this has not been predicted. We hypothesize that it may be a statistical fluctuation consistent with the overall noise fluctuations and speculate such fluctuations drive the differential lag between the superfluid and the crust above its critical value, thus triggering the glitch. This suggests a large number of glitches could be preceded by a slow-down, providing testable predictions.

Analyses like that presented herein only assess the relative evidence of models. We focus on phenomenological, albeit physically-motivated models, in a bid to remain model agnostic. Even the best fitting models tested here do not explain all the features in the data, e.g., Fig. 3. Future explorations may uncover new descriptions that explain the data better than the models used herein. For example, further theoretical modelling may provide a more nuanced view of how the slow-down preceding the glitch should manifest; the method we developed is easily extendable to compare more complex models.

While direct modelling is one avenue of further investigation, model-agnostic approaches may also yield considerable insight. Figure 3 is a first step in this direction, although it has the subtlety that the time window distorts temporal and amplitude features. Another method could be e.g., shapelet-based models for the frequency evolution, providing a means to study the underlying frequency evolution without modelling constraints.

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Contributions to the paper G.A. is responsible for the data analysis; G.A., P.D.L., and V.G. are responsible for the model development and discussion; J.P. is responsible for the data collection and reduction.

Competing Interests Authors declare no competing interests.

METHODS

We use the 72-min stretch of data collected by the University of Tasmania Mt Pleasant 26-m radio telescope² on 2016 December 12. The raw flux is analysed fitting a standard pulse template to individual pulses and estimating the site arrival time of each pulse. We use `Tempo2`^{17,18} to convert site arrival times to solar-system barycentre times, and use the Bayesian analysis package `Bilby`¹⁹ to fit timing models.

The likelihood for the i th pulse with observed arrival time t_i is calculated from

$$\mathcal{L}(t_i; \theta) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{(t_i - h(i; \theta))^2}{2\sigma^2}\right], \quad (4)$$

where $h(i; \theta)$ is the predicted arrival time of the i th pulse (within the context of the model). The variance of this distribution is further given by $\sigma^2 = \sigma_i^2 + \sigma_0^2$, where σ_i^2 is the estimated variance of the i th arrival time (as output by the matched-filter profile analysis) and σ_0 is an additional stochastic noise to be fit for. The priors used are listed in Table 2.

	Prior distribution	Units	Models
f_0	Uniform(11.1854, 11.1874)	Hz	all
ϕ_0	Uniform(-5, 5)	—	all
σ_0	Uniform(0, 0.01)	s	all
Δf	Log-Uniform(10^{-8} , 10^{-4})	Hz	all
t_g	Uniform(-100, 100)	s	all
τ_r	Uniform(0, 1000)	s	H_1, H_2, H_{2+p}
Δf_d	Log-Uniform(10^{-8} , 10^{-4})	Hz	H_2, H_{2+p}
τ_d	Uniform(0, 1000)	s	H_2, H_{2+p}
Δt	Uniform(0, 500)	s	H_{2+p}
Δf_p	Half-Normal(0, 10^{-5})	Hz	H_{2+p}

Table 2 | Table of priors used throughout this work. For the parameters not introduced in the text, ϕ_0 is the phase parameter (number of rotations), we provide a wider prior to allow the reference pulse to not be zero; and σ is the standard-deviation of the Gaussian likelihood.

Data availability The data used in this work is available from Ref.².

Code availability The `bilby`¹⁹ analysis code is available from <https://git.ligo.org/lscsoft/bilby> and particular scripts for this analysis are available on request from the authors.